

The Saving-Investment Relationships: A Markov Switching Causality Analysis

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Abstract:

This paper investigates the empirical saving-investment relationships for Côte d'Ivoire and Ghana, using a Markov Switching VAR model. We find regime-dependent causality from saving to investment in Côte d'Ivoire but not in Ghana. In terms of Feldstein and Horioka (1980) capital mobility hypothesis these findings suggest a more capital mobility in Ghana than in Côte d'Ivoire implying that foreign capital flows towards Côte d'Ivoire and not to Ghana during the studied period.

JEL Classification: I20, C22, C51, O54

Keywords: Saving, Investment, Markov Switching VAR, Causality.

1. Introduction

Following the seminal paper by Feldstein and Horioka (1980), the relationships between saving and investment have been studied by several authors including Baxter and Crucini (1993); Coaley and Kulazsi, (1997); Obstfeld and Rogoff (1995); Cadoret (2001). The suggested relationship between saving and investment by Feldstein and Horioka (1980) is modelled as follows:

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$$\left(\frac{I}{Y}\right)_{it} = \alpha + \beta\left(\frac{S}{Y}\right)_{it} + \varepsilon_{it} \quad (1)$$

where (I/Y) is the ratio of gross domestic investment to domestic product, (S/Y) is the ratio of gross domestic saving to domestic product; β is the saving-investment retention coefficient and an error term.

This equation (1) allows for the investigation of the full capital mobility hypothesis. In case of perfect capital mobility there should be no relation between domestic saving and investment and β is expected to be close to zero, which means that investment decision in one country should be independent of domestic level of saving and reciprocally. The interpretation is that saving in each country responds globally to worldwide opportunities for higher profit investment while domestic investment is financed by the worldwide capital.

While a lot of literature has extensively analysed the Feldstein and Horioka equation; the results obtained are mixed according to time period. These studies make use of various methods as cross-section or panel data model estimation (Feldstein and Horioka (1980), Feldstein (1983)); time series modeling (Obstfeld (1986), Seshaiyah and Sriyval (2005)), to our knowledge none has used a non-linear model, particularly for developing countries. In this paper we aim at filling this gap by extending the existing literature on this matter, with a special emphasis on two African countries' experiences.

This paper is an attempt to use an alternative methodology based on multivariate Markov switching model to test for causal link between saving and investment. Two reasons motivate the use of such a model: (i) the worldwide environment is fluctuating and thus one could expect the value of the β parameter to be non-constant across time span. Cadoret (2001) has confirmed this observation and to overcome the variability of the coefficient, the equation (1) has been often estimated

using subperiods after detecting structural breaks on the entire period (Kanas, 2005). This procedure supposes a prior knowledge of the break dates, (ii) splitting the sample empowers the data and doesn't allow to seizing the whole phenomenon. Moreover breaks are recurrent and require non-linear model as Markov switching model.

The main objective of the paper is to use a multivariate Markov switching model introduced by Krolzig (1997a; 1997b), Krolzig and Toro (1999) as a generalization of Hamilton (1989; 1990) univariate model to investigate the relationships between saving and investment. The MS-VAR approach has the advantage not only to avoid splitting the sample period under study into subperiods but also the variability and structural change of the link between saving and investment over time is endogenously taken into account in the model.

Our major findings are as follows. We find regime-dependent causality from saving to investment in Côte d'Ivoire but not in Ghana.

Section 2 presents the methodology used in the paper, Section 3 the econometric results and Section 4 concludes.

2. Econometric Methodology

Suppose we intend to analyse Granger causality between the bivariate series $\{x_t' = [x_{1,t} : x_{2,t}]\}$ Granger causality analysis is based on the following MS-VAR model:

$$\Delta x_{1,t} = \mu_{x_1}(s_t) + \sum_{k=1}^p \phi_{x_1, x_1, k}(s_t) \Delta x_{1,t-1} + \sum_{k=1}^p \phi_{x_1, x_2, k}(s_t) \Delta x_{2,t-k} + Z(s_t) u_{x_1,t}$$

$$u_{x_1,t} \quad \text{and} \quad u_{x_2,t} \sim \text{nid}(0,1)$$

$$\Delta x_{2,t} = \mu_{x_2}(s_t) + \sum_{k=1}^p \phi_{x_2, x_1, k}(s_t) \Delta x_{1,t-1} + \sum_{k=1}^p \phi_{x_2, x_2, k}(s_t) \Delta x_{2,t-k} + Z(s_t) u_{x_2,t}$$
(1)

where s_t is an unobservable random variable indicating the state of regime at date t , and $Z(s_t)$ is a regime dependent matrix. The transition probabilities p_{ij} are assumed:

$$p_{ij} = P(s_{t+1} = j | s_t = i), \quad \sum_{j=1}^2 p_{ij} = 1 \quad \forall i, j \in \{1,2\} \quad (2)$$

To test for Granger causality from x_2 to x_1 in regime 1, the null hypothesis of non-causality is

$$H_0 : \phi_{x_1, x_2, 1}(s_t = 1) = \dots = \phi_{x_1, x_2, p}(s_t = 1) = 0 \quad (3)$$

This null hypothesis test is conducted using the MS-VAR model by imposing restriction on the values of the autoregressive parameters.

In fact the significance of the regime dependent autoregressive parameter ϕ_{x_1, x_2} in equation 1 infers causality from x_2 to x_1 in regime 1. On the other hand the significance of ϕ_{x_2, x_1} infers causality from x_1 to x_2 in the concerned regime.

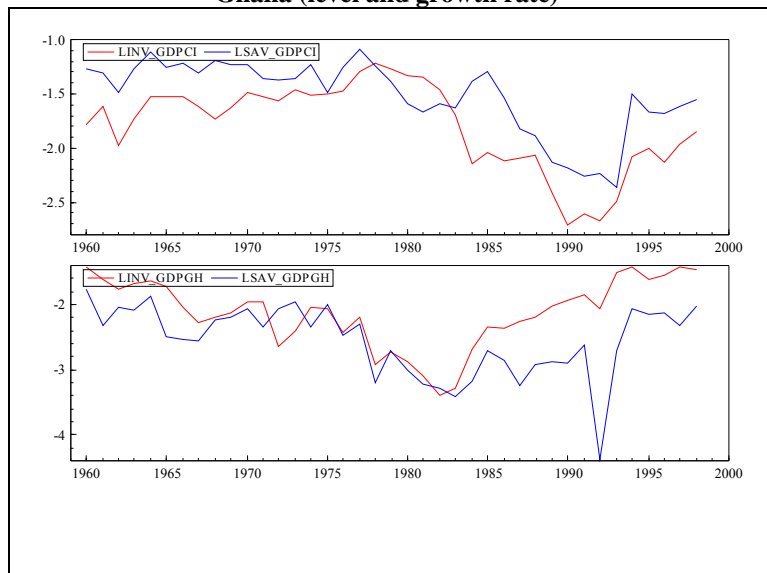
3. Econometric Results

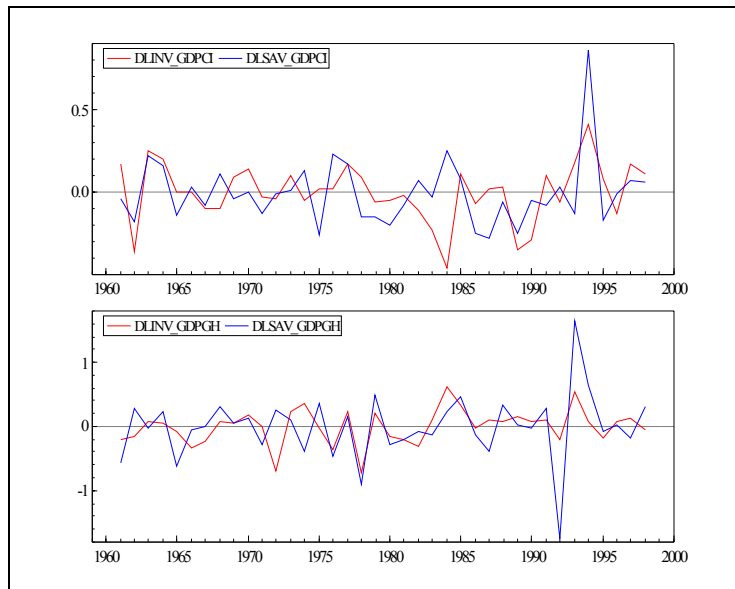
The yearly data for Côte d'Ivoire and Ghana are from World Development Indicators (WDI) and cover the period

1960–1998. We choose 2 neighboring countries Côte d’Ivoire, which belong to a monetary union (WAEMU) and Ghana, which has own money. Investment is measure by (I/Y) the ratio of gross domestic investment to domestic product, while saving (S/Y) is the ratio of gross domestic saving to domestic product. Figure 1 contains the log of the variable (top panel) and the growth rate (bottom panel). We observe that saving and investment grow from 1960 to the peak around 1979 in Côte d’Ivoire and decline to the trough around 1990 for investment and around 1993 for saving. In average saving has been higher than investment in Côte d’Ivoire In Ghana saving and investment decrease from 1960 to the trough around 1982-1983 and after 1982, investment grows faster than saving, investment being higher than saving in average.

It could be observed from Figure 1 that saving could not finance investment in Ghana but the reverse holds in Côte d’Ivoire.

Figure 1: Saving and investment to GDP in Côte d’Ivoire and Ghana (level and growth rate)





The variable $x_t' = [inv_t = x_{1,t} : sav_t = x_{2,t}]$ and the estimations results are provided in Table 1². The likelihood ratio (LR) test suggests that the null hypothesis of no regime switching or equivalently to linear VAR model against the alternative of regime switching is rejected. The rejection is equivalent to the rejection of linear VAR in favour of the MS-VAR model. Moreover the log likelihood (LL) indicates that a 2-Lag VAR (against 1 Lag) is suitable to the MS-VAR in the Côte d'Ivoire while a 1-Lag VAR is suitable in Ghana, which is confirmed by the computed Bayes Factor (BF)³. These results

² Estimations are realized using MS-VAR package from PcGive 10 (Hendry and Doornik 2001).

³The BF(1;2) interpretation (see Poirier 1995):

B12 > 1, evidence supports H1

10e-1/2 < B12 < 1, very slight evidence against H1

10e-1 < B12 < 10e-1/2, slight evidence against H1

10e-2 < B12 < 10e-1, strong to very strong evidence against H1

B12 < 10e-2, decisive evidence against H1

indicate that the bivariate saving investment relationship is characterized by volatility regime switching and subject to regime switching.

The standard deviation of variables across regime is 0.10 in Côte d'Ivoire for saving and 0.13 for invest in regime 1 while in regime 2, standard deviation is 0.10 for saving and 0.05 for investment. In Côte d'Ivoire Investment and saving are both more volatile in regime 1 than in regime 2.

In Ghana investment and saving are both more volatile in regime 1 than in regime 2 (standard deviation are 0.39 for saving in regime 1 and 0.22 in regime 2, while standard deviation of invest is 0.32 in regime 1 and only 0.10 in regime 2). Regime 2 is a low volatility regime in Ghana and in Côte d'Ivoire as well.

The transition probability from regime 1 to regime 1 is 0.82 and the transition probability from regime 2 to regime 2 is 0.35, which indicates that regime 1 is very persistent compare to regime 2 in Côte d'Ivoire. But in Ghana the two regimes are persistent (0.78 in regime 1 and 0.77 in regime 2).

The value of the contemporaneous correlation reflects the contemporaneous link between saving and investment. During regime 1 (the high volatility regime) the correlation is negative -1.167 and positive 0.904 in the low volatility regime (regime 2) for Côte d'Ivoire. In Ghana this correlation is 0.366 in regime 1 and 0.086 in regime 2 thus positive in both regimes.

These results reveal a dynamic regime dependent link between saving and investment both in Côte d'Ivoire and Ghana. In Côte d'Ivoire the correlation is high during low volatility regime while in Ghana the correlation is high in the high volatility regime.

We notice that the regime dependent autoregressive parameters $\phi_{sav,inv,1}$ and $\phi_{sav,inv,2}$ are statistically significant in regime 2 but not in regime 1 for Côte d'Ivoire. These findings suggest that investment has caused saving only in regime 2 in Côte d'Ivoire.

Similarly we can see that $\phi_{inv,sav,2}$ is statistically significant in both regimes while $\phi_{inv,sav,1}$ is significant only on regime 1, meaning that causality runs from saving to investment in Côte d'Ivoire in regime 1. None of the autoregressive parameters are significant in Ghana meaning that causality from saving to investment is rejected in Ghana.

To formally test for causality from saving to investment we restrict the model and test the null of the autoregressive parameters being zero (the non causality hypothesis). The Likelihood Ratio of the unrestricted model is always greater than the restricted one. The LR-test results (Table 2) confirm the findings for Côte d'Ivoire and Ghana. We find causality from saving to investment in Côte d'Ivoire but not in Ghana.

In terms of capital mobility these findings could be interpreted as a more capital mobility in Ghana where saving and investment are independent than in Côte d'Ivoire where they are not. This implies that foreign capital flows towards Côte d'Ivoire and not to Ghana in the studied period, which could be explained in part by the fact that Côte d'Ivoire being embedded in monetary zone (WAEMU) with a supra national currency makes this country a more secure place for investors than Ghana which has its own currency. In effect Côte d'Ivoire was a very stable country for investors before the political coup of December 1999 and war since September 2002, while Ghana has been unstable from 1966 until recent years. Moreover the investment code (law) in Côte d'Ivoire is very favorable to

investors, which are allowed to take a large part of earnings out of the country.

Table 1: MS-VAR Model Estimation

<i>Parameters</i>	Côte d'Ivoire MSIAH(2)-VAR(2)		Ghana MSIAH(2)-VAR(1)	
	<i>Regime 1</i>	<i>Regime 2</i>	<i>Regime 1</i>	<i>Regime 2</i>
μ_{sav}	-0.035 (-1.68)	0.277* (5.33)	-0.255* (-2.53)	0.058 (0.79)
$\phi_{sav,sav,1}$	-0.146 (-0.93)	-1.745* (-7.64)	-0.937* (-4.82)	0.024 (0.15)
$\phi_{sav,sav,2}$	0.378* (2.39)	-1.087 (*-6.06)		
$\phi_{sav,inv,1}$	0.211 (1.09)	-0.651* (-5.62)	-0.195 (-1.25)	-0.085 (-0.96)
$\phi_{sav,inv,2}$	0.110 (0.460)	-0.654* (-7.21)		
μ_{inv}	-0.020 (-0.74)	0.196* (8.05)	-0.092 (-1.12)	0.042 (1.004)
$\phi_{inv,sav,1}$	-0.080 (-0.59)	1.838* (6.15)	-0.225 (-0.64)	0.454 (1.49)
$\phi_{inv,sav,2}$	-0.743* (-4.24)	0.624* (3.92)		
$\phi_{inv,inv,1}$	0.316 (1.82)	0.654* (4.45)	-0.183 (-0.65)	0.317* (1.96)
$\phi_{inv,inv,2}$	-0.172 (-0.78)	0.280* (3.36)		
σ_{sav}	0.1035	0.1029	0.399	0.228
σ_{inv}	0.1345	0.0549	0.326	0.106
Contemporaneous Correlation	-1.167	0.904	0.366	0.086
p_{11}		0.820		0.786
p_{22}		0.355		0.772
log Like.		50.003 (25.48)		-10.858 (-23.56)
LR		49.04		25.40
AIC		-1.22 (-0.69)		1.66 (1.76)
HQ		-0.79 (-0.49)		1.97 (1.89)
Bayes Factor (BF)		BF(1;2) = 37.10 / 50.00 = 0.60		BF(1;2) = -10.85 / -4.56 = 2.37

Notes: t-ratios in parentheses; * denotes significance at 0.05 level; BF(1;2) is Bayes factor 1-Lag model against 2-Lag model.

Table 2: Causality Test

Model	Côte d'Ivoire	Ghana
	LR	LR
Unrestricted	50.00	-10.85
Restricted	41.30	-12.62
LR-test	17.39	3.5
χ^2_8	15.5	15.5
Result	Reject the null	Accept the null

5. Conclusion

The variability and structural change of the link between saving and investment over time is endogenously taken into account in this paper. We investigate the saving-investment relationships for two African countries using a MS-VAR approach, which reveals the non-homogeneity of the relation between the 2 variables over the period of study. We find causality from saving to investment in Côte d'Ivoire but not in Ghana.

In terms of capital mobility these findings suggest a more capital mobility in Ghana than in Côte d'Ivoire implying that foreign capital flows towards Côte d'Ivoire and not to Ghana. With the history of both countries in mind, Côte d'Ivoire being very stable politically before the coup (December 1999) and war (September 2002) while Ghana was instable since 1966 up to recent years, these results confirm the prior need for a peaceful environment to attract foreign investment in the sub-Saharan African region.

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